

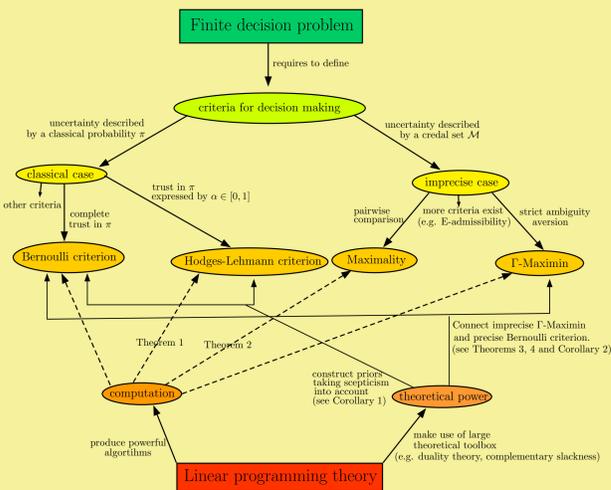
# Decision theory meets linear optimization beyond computation

Christoph Jansen & Thomas Augustin, Department of Statistics, University of Munich

Contact: christoph.jansen@stat.uni-muenchen.de augustin@stat.uni-muenchen.de



## 1. Guideline through the poster



## 2. Background: Some finite decision theory

We consider the standard model of *finite* decision theory:

- $\mathbb{A} = \{a_1, \dots, a_n\}$ ,  $n \in \mathbb{N}$  (set of actions)
- $\Theta = \{\theta_1, \dots, \theta_m\}$ ,  $m \in \mathbb{N}$  (states of nature)
- $u : \mathbb{A} \times \Theta \rightarrow \mathbb{R}$  (cardinal utility function)

Naturally, the utility function associates

- every action  $a \in \mathbb{A}$  with a gamble  $u_a$  on  $(\Theta, 2^\Theta)$ :  

$$u_a : \Theta \rightarrow \mathbb{R}, \quad \theta \mapsto u(a, \theta) \quad (1)$$
- every state  $\theta \in \mathbb{A}$  with a variable  $u^\theta$  on  $(\mathbb{A}, 2^{\mathbb{A}})$ :  

$$u^\theta : \mathbb{A} \rightarrow \mathbb{R}, \quad a \mapsto u(a, \theta) \quad (2)$$

With  $u_{ij} := u(a_i, \theta_j)$ , we represent the model by its *utility matrix*:

	$\theta_1$	...	$\theta_m$	
$a_1$	$u_{11}$	...	$u_{1m}$	$u_{a_i} : \Theta \rightarrow \mathbb{R}$ associated gambles
$a_2$	$u_{21}$	...	$u_{2m}$	
$\vdots$	$\vdots$	...	$\vdots$	
$a_n$	$u_{n1}$	...	$u_{nm}$	

associated random variables

Depending on the context, we also allow for choosing *randomized actions*, i.e. classical probability measures on  $(\mathbb{A}, 2^{\mathbb{A}})$ . We denote the set of all randomized actions by  $G(\mathbb{A})$ .

The utility function  $u$  is then extended to a utility function  $G(u)$  on  $G(\mathbb{A}) \times \Theta$  by assigning each pair  $(\lambda, \theta)$  the expectation of the random variable  $u^\theta$  under the measure  $\lambda$ , i.e.  $\mathbb{E}_\lambda[u^\theta]$ .

Every *pure* action  $a \in \mathbb{A}$  then can uniquely be identified with the *Dirac-measure*  $\delta_a \in G(\mathbb{A})$  and we have  $u(a, \theta) = G(u)(\delta_a, \theta)$  for all  $(a, \theta) \in \mathbb{A} \times \Theta$ . Further, also (1) can easily be extended to randomized actions by defining, for every  $\lambda \in G(\mathbb{A})$  fixed,  $G(u)_\lambda(\theta) := G(u)(\lambda, \theta)$  for all  $\theta \in \Theta$ .

## 3. A criterion from classical decision theory

Apart from the border cases of *maximizing expected utility* w.r.t. a precise prior and the *maximin-criterion*, classical decision theory tries to cope with decision making under vague information, too: The criterion of *Hodges and Lehmann* allows the decision maker to model his *degree of trust* in the prior by a parameter  $\alpha \in [0, 1]$ .

Specifically, if  $\pi$  is a probability measure on  $(\Theta, 2^\Theta)$ , a randomized action  $\lambda^* \in G(\mathbb{A})$  is said to be *Hodges-Lehmann-optimal* w.r.t.  $\pi$  and  $\alpha$  (short:  $\Phi_{\pi, \alpha}$ -optimal), if  $\Phi_{\pi, \alpha}(\lambda^*) \geq \Phi_{\pi, \alpha}(\lambda)$  for all  $\lambda \in G(\mathbb{A})$ , where

$$\Phi_{\pi, \alpha}(\lambda) := (1 - \alpha) \cdot \min_{\theta} G(u)(\lambda, \theta) + \alpha \cdot \mathbb{E}_\pi[G(u)_\lambda] \quad (3)$$

Theorem 1 describes an algorithm determining a randomized Hodges-Lehmann-actions for arbitrary pairs  $(\pi, \alpha)$ .

**Theorem 1.** Consider the linear programming problem

$$(1 - \alpha) \cdot (w_1 - w_2) + \alpha \cdot \sum_{i=1}^n \mathbb{E}_\pi(u_{a_i}) \cdot p_i \rightarrow \max_{(w_1, w_2, p_1, \dots, p_n)} \quad (4)$$

with constraints  $(w_1, w_2, p_1, \dots, p_n) \geq 0$  and

- $\sum_{i=1}^n p_i = 1$
- $w_1 - w_2 \leq \sum_{i=1}^n u_{ij} \cdot p_i$  for all  $j = 1, \dots, m$ .

Then the following holds:

- Every optimal solution  $(w_1^*, w_2^*, p_1^*, \dots, p_n^*)$  to (4) induces a  $\Phi_{\pi, \alpha}$ -optimal randomized action  $\lambda^* \in G(\mathbb{A})$  by setting  $\lambda^*(\{a_i\}) := p_i^*$ .
- There always exists a  $\Phi_{\pi, \alpha}$ -optimal randomized action.

By applying *duality* theory, we receive the following Corollary. Its proof can be interpreted as a method to construct priors that take the actor's *scepticism about  $\pi$*  (expressed by  $\alpha$ ) into account.

**Corollary 1.** Let  $\lambda^* \in G(\mathbb{A})$  denote a  $\Phi_{\pi, \alpha}$ -optimal randomized action. Then, there exists a probability measure  $\mu_{\pi, \alpha}$  on  $(\Theta, 2^\Theta)$  and a pure action  $a^* \in \mathbb{A}$  such that

$$\Phi_{\pi, \alpha}(\lambda^*) = \mathbb{E}_{\mu_{\pi, \alpha}}[u_{a^*}] \quad (5)$$

## 4. Linear partial information

Kofler and Menges' theory of *linear partial information* (see [4]) assumes the uncertainty underlying the decision situation to be expressible by a *convex credal set*  $\mathcal{M}$  on  $(\Theta, 2^\Theta)$  of the form

$$\mathcal{M} := \{\pi \mid \underline{b}_s \leq \mathbb{E}_\pi(f_s) \leq \bar{b}_s \forall s = 1, \dots, r\} \quad (6)$$

where, for all  $s = 1, \dots, r$ , we have  $(\underline{b}_s, \bar{b}_s) \in \mathbb{R}^2$  such that  $\underline{b}_s \leq \bar{b}_s$  and  $f_s : \Theta \rightarrow \mathbb{R}$ . Note that these sets correspond to the credal sets induced by finite sets of gambles  $\mathcal{K}$  from Walley's theory.

Here, criteria for decision making strongly depend on the actor's *attitude towards ambiguity*, i.e. the non-stochastic uncertainty between the measures contained in  $\mathcal{M}$ . Accordingly, many concurring criteria exist (see for instance [3]). Linear programming based results for a selection of them are presented in the following sections.

## 5. Checking maximality of pure actions

An action  $a^* \in \mathbb{A}$  is said to be  *$\mathcal{M}$ -maximal*, if

$$\forall a \in \mathbb{A} \exists \pi_a \in \mathcal{M} : \mathbb{E}_{\pi_a}(u_{a^*}) \geq \mathbb{E}_{\pi_a}(u_a) \quad (7)$$

Naturally, the above definition extends to randomized actions. For randomized actions,  $\mathcal{M}$ -maximality and  $E(\mathcal{M})$ -admissibility coincide. A algorithm for determining the set of all randomized  $E(\mathcal{M})$ -admissible actions has been introduced in [1, section 5.2].

However, for finite  $\mathbb{A}$ , being  $\mathcal{M}$ -maximal is a strictly weaker condition and, therefore, needs to be checked separately from  $E(\mathcal{M})$ -admissibility. Theorem 2 describes a linear programming based algorithm for checking  $\mathcal{M}$ -maximality of a pure  $a^* \in \mathbb{A}$ .

**Theorem 2.** Let  $(\mathbb{A}, \Theta, u(\cdot))$  denote a finite decision problem and let  $\mathcal{M}$  be of the form (6). Consider the linear program

$$\sum_{i=1}^n \left( \sum_{j=1}^m i \gamma_j \right) \rightarrow \max_{(\gamma_1, \dots, \gamma_m)} \quad (8)$$

with constraints  $(\gamma_1, \dots, \gamma_m) \geq 0$  and

- $\sum_{j=1}^m i \gamma_j \leq 1$  for all  $i = 1, \dots, n$
- $\underline{b}_s \leq \sum_{j=1}^m f_s(\theta_j) \cdot i \gamma_j \leq \bar{b}_s$  for all  $s = 1, \dots, r$ ,  $i = 1, \dots, n$
- $\sum_{j=1}^m (u_{ij} - u_{*j}) \cdot i \gamma_j \leq 0$  for all  $i = 1, \dots, n$

Then  $a^* \approx (u_{*1}, \dots, u_{*m}) \in \mathbb{A}$  is  $\mathcal{M}$ -maximal iff the optimal outcome of (8) equals  $n$ .

## 6. $\Gamma$ -Maximin and least favourable priors

For a probability measure  $\pi$  on  $(\Theta, 2^\Theta)$ , let  $B(\pi)$  denote the Bayes-utility w.r.t.  $\pi$  (that is  $B(\pi) = \mathbb{E}_\pi(u_{a^*})$ , where  $a^* \in \mathbb{A}$  denotes an arbitrary Bayes-action w.r.t.  $\pi$ ). The set of all Bayes-actions w.r.t.  $\pi$  is denoted by  $\mathbb{A}_\pi$ .

If  $\mathcal{M}$  is a credal set of the form defined in (6), we call  $\pi^- \in \mathcal{M}$  a *least favourable prior (lfp)* from  $\mathcal{M}$  iff  $B(\pi^-) \leq B(\pi)$  holds for all  $\pi \in \mathcal{M}$ . Theorem 3 describes a linear programming approach for determining a least favourable prior from  $\mathcal{M}$ .

**Theorem 3.** Let  $(\mathbb{A}, \Theta, u(\cdot))$  denote a decision problem and let  $\mathcal{M}$  be of the form (6). Consider the linear program

$$w_1 - w_2 \rightarrow \min_{(w_1, w_2, \pi_1, \dots, \pi_m)} \quad (9)$$

with constraints  $(w_1, w_2, \pi_1, \dots, \pi_m) \geq 0$  and

- $\sum_{j=1}^m \pi_j = 1$
- $\underline{b}_s \leq \sum_{j=1}^m f_s(\theta_j) \cdot \pi_j \leq \bar{b}_s$  for all  $s = 1, \dots, r$
- $w_1 - w_2 \geq \sum_{j=1}^m u_{ij} \cdot \pi_j$  for all  $i = 1, \dots, n$

Then the following holds:

- Every optimal solution  $(w_1^*, \dots, w_m^*)$  to (9) induces a least favourable prior  $\pi^- \in \mathcal{M}$  by setting  $\pi^-(\{\theta_j\}) := \pi_j^*$ .
- There always exists a least favourable prior.

Next, we show some connections between least favourable priors and randomized  $\Gamma$ -Maximin actions. We start by recalling the  $\Gamma$ -Maximin criterion: A randomized action  $\lambda^* \in G(\mathbb{A})$  is said to be  *$\mathcal{M}$ -Maximin*

optimal iff for all  $\lambda \in G(\mathbb{A})$ :

$$\min_{\pi \in \mathcal{M}} \mathbb{E}_\pi[G(u)_\lambda] \geq \min_{\pi \in \mathcal{M}} \mathbb{E}_\pi[G(u)_\lambda] \quad (10)$$

It turns out that the linear program from Theorem 3 is *dual* to the one for determining a  $\mathcal{M}$ -Maximin optimal randomized action described in [1, section 3.2]. Together with the *complementary slackness property* from linear optimization theory, this allows to derive deep connections between least favourable priors and the  $\Gamma$ -Maximin criterion.

**Theorem 4.** Let  $(\mathbb{A}, \Theta, u(\cdot))$  denote a finite decision problem and let  $\mathcal{M}$  be of the form (6). Then the following holds:

- If  $\pi^-$  is a lfp from  $\mathcal{M}$ , then for all optimal randomized  $\mathcal{M}$ -Maximin actions  $\lambda^* \in G(\mathbb{A})$  we have  $\lambda^*(\{a\}) = 0$  for all  $a \in \mathbb{A} \setminus \mathbb{A}_{\pi^-}$ .
- Let  $\lambda^* \in G(\mathbb{A})$  be an optimal randomized  $\mathcal{M}$ -Maximin action. If, for  $a \in \mathbb{A}$ , we have  $\lambda^*(\{a\}) > 0$ , then  $a \in \mathbb{A}_{\pi^-}$  for all least favourable priors  $\pi^-$  from  $\mathcal{M}$ .
- Let  $\pi^-$  denote a lfp from  $\mathcal{M}$  and let  $\lambda^* \in G(\mathbb{A})$  denote a randomized  $\mathcal{M}$ -Maximin action. Then for all  $a \in \mathbb{A}_{\pi^-}$  we have

$$\mathbb{E}_{\pi^-}[u_a] = \mathbb{E}_{\mathcal{M}}[G(u)_\lambda^*]$$

As an immediate consequence of Theorem 4, we can specify conditions under which randomization cannot improve utility, if optimality is defined in terms of the  $\Gamma$ -maximin criterion.

**Corollary 2.** If there exists a least favourable prior  $\pi^-$  from  $\mathcal{M}$  such that  $\mathbb{A}_{\pi^-} = \{a_z\}$  for some  $z \in \{1, \dots, n\}$ , then  $\delta_{a_z} \in G(\mathbb{A})$  is the unique randomized  $\mathcal{M}$ -Maximin action. Particularly, randomization is unnecessary in such situations.

## 7. A toy example

Consider the decision problem given by the table

$u_{ij}$	$\theta_1$	$\theta_2$	$\theta_3$	$\theta_4$
$a_1$	20	15	10	5
$a_2$	30	10	10	20
$a_3$	20	40	0	20

and assume that uncertainty is described by the credal set

$$\mathcal{M} := \{\pi \mid 0.3 \leq \pi_2 + \pi_3 \leq 0.7\}$$

- **Section 6:** Applying the algorithm from Theorem 3 gives the optimal solution  $(13, 0, 0, 0, 0.7, 0.3)$ . Thus, a least favourable prior  $\pi^-$  from  $\mathcal{M}$  is induced by the vector  $(0, 0.7, 0.3, 0)$ . Simple computation gives  $\mathbb{A}_{\pi^-} = \{a_2\}$ . Therefore, according to Corollary 2,  $a_2$  is the unique  $\mathcal{M}$ -Maximin action (even compared to randomized actions) with utility 13.
- **Section 5:** Resolving the linear programming problem from Theorem 2 for actions  $a_1, a_2$  and  $a_3$  gives optimal value 3 for each of them. Thus, all available actions are  $\mathcal{M}$ -maximal.
- **Section 3:** Let  $\tau$  denote the prior on  $(\Theta, 2^\Theta)$  induced by  $(0.2, 0.7, 0.05, 0.05)$  and let our trust in  $\tau$  be expressed by  $\alpha = 0.3$ . Resolving the linear programming problem from Theorem 1 then gives the optimal solution  $(8, 0, 0.8, 0, 0.2)$ . Thus, a  $\Phi_{\tau, 0.3}$ -optimal randomized action  $\lambda^* \in G(\mathbb{A})$  is induced by  $(0.8, 0, 0.2)$ .

Next, we can use the *constructive* proof of Corollary 1 to compute the measure  $\mu_{\tau, 0.3}$  on  $(\Theta, 2^\Theta)$  defined in Corollary 1. The measure  $\mu_{\tau, 0.3}$  is induced by the vector  $(0.070, 0.245, 0.656, 0.029)$ .

**Implementation:** The R-code for the toy example is available on <http://www.statistik.lmu.de/~cjansen/index.html>

## Outlook: Future research

Investigating further consequences of Theorem 4: What can we learn by restricting the set  $\mathcal{M}$  to special cases (for instance *comparative probability* or *non-degenerated credal sets*)?

## References

- [1] L.V. Utkin, T. Augustin. Powerful algorithms for decision making under partial prior information and general ambiguity attitudes. In: F.G. Cozman, R. Nau, T. Seidenfeld (eds.): *ISIPTA '05*, 2005, pp. 349-358.
- [2] D. Kikuti, F.G. Cozman, C.P. de Campos. Partially ordered preferences in decision trees: computing strategies with imprecision in probabilities. In: R. Brafman, U. Junker (eds.): *Multidisciplinary IJCAI-05 Workshop on Advances in Preference Handling*, 2005, pp. 118-123.
- [3] N. Huntley, R. Hable, M.C.M. Troffaes. Decision making. In: *Introduction to imprecise probabilities*. Ed. by T. Augustin, F.P.A. Coolen, G. de Cooman, M.C.M. Troffaes. Chichester: Wiley, 2014, pp. 190-206.
- [4] E. Kofler, G. Menges. *Entscheiden bei unvollständiger Information*. Springer, Berlin (Lecture Notes in Economics and Mathematical Systems, 136), 1976.
- [5] Hodges, Joseph L. and Lehmann, Erich L. The use of Previous Experience in Reaching Statistical Decisions. In: *The Annals of Mathematical Statistics* 23.3 (1952), pp. 396407.