Wednesday 14:00-17:30

Part 7 Bayesian hierarchical modelling, simulation and MCMC

by Gero Walter

Bayesian hierarchical modelling, simulation and MCMC _{Outline}

Bayesian hierarchical modelling / Bayesian networks / graphical models

Exercises I

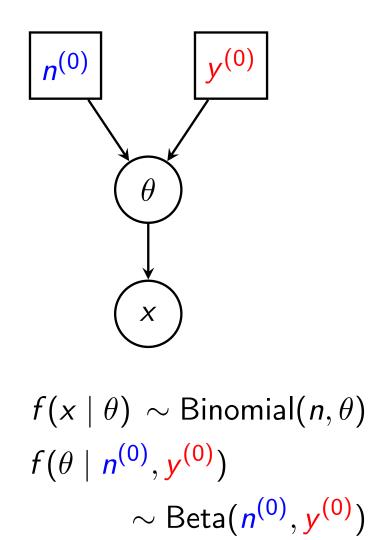
Simulation & MCMC

Exercises II

Bayesian Hierarchical Modelling, a.k.a. Bayesian (Belief) Networks, a.k.a. Graphical Models

- many names for the same thing (it's a powerful tool),
 I will use the term Bayesian Networks (BNs)
- BNs as a unifying way to think about (Bayesian) statistical models
- how to build complex Bayesian models out of simple building blocks
- how to specify joint distributions (over many variables) via univariate distributions using conditional independence assumptions
- conditional independence assumptions are visualized by a graph
- the graph can establish a hierarchy between variables

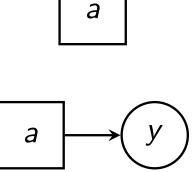
Bayesian Networks: Simple Example



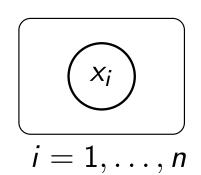
variables / parameters (fixed values

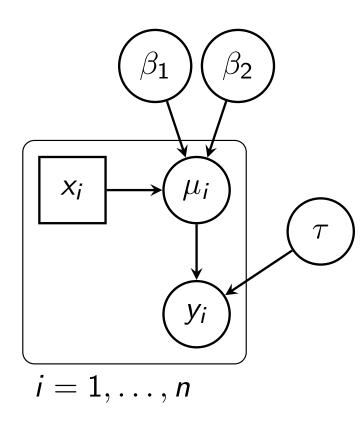
y depends on a

repeated nodes



У





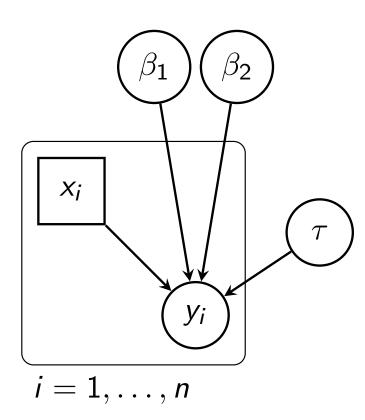
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$$\tau \mid a, b \sim \text{Gamma}(a, b)$$

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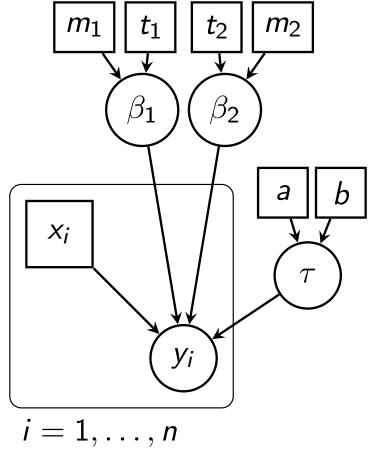
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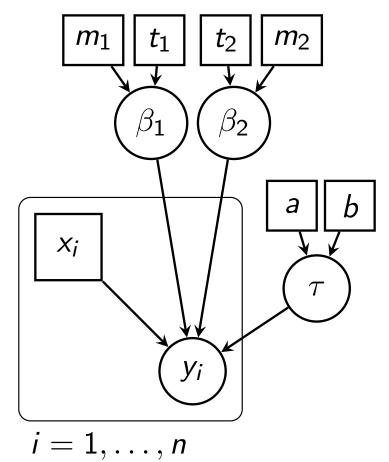
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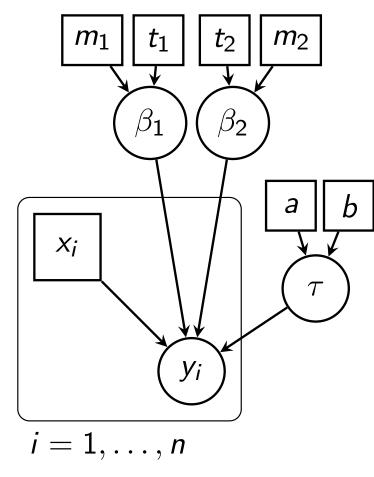
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- BNs enable us to construct probability distributions that capture the important dependencies between the relevant variables in a given inference problem while keeping models (relatively) simple
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- What kind of graphs work for expressing conditional independence relations?

Bayesian Networks: Directed Acyclic Graphs

Definition (Directed Graph)

A directed graph G = (V, E) consists of a set of vertices V and a set of edges E, where $E \subset V \times V$. An arrow leads from $u \in V$ to $v \in V$ if and only if $(u, v) \in E$; u is the source and v is the target of edge (u, v).

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A path in a graph is an ordered set of edges $\{e_i\}$ such that $t(e_i) = s(e_{i+1})$ (chain of head-to-tail arrows). A cycle is a path such that $t(e_N) = s(e_1)$, where N is the number of edges in the path.

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Definition (Directed Acyclic Graph)

A *directed acyclic graph (DAG)* is a directed graph that does not contain a cycle, i.e. there does not exist a subset of edges that forms a cycle.

Bayesian Networks: Formal Definition

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Given a DAG G = (V, E), and variables $x_V = \{x_v\}_{v \in V}$, a *Bayesian network* with respect to G and x_V is a joint probability distribution for the x_V of the form:

$$f(x_V) = \prod_{v \in V} f(x_v \mid x_{\mathsf{pa}(v)})$$

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joint distribution factorizes according to the graph!

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$$f(x_1,...,x_K) = f(x_K \mid x_1,...,x_{K-1}) f(x_{K-1} \mid x_1,...,x_{K-2})$$

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corresponds to a *fully connected graph*: there is a link between every pair of vertices (each of the K vertices has incoming edges from all lower-numbered vertices)

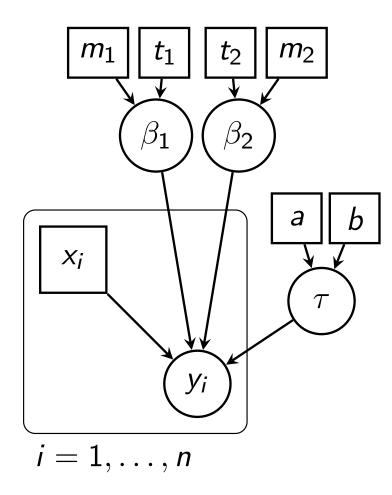
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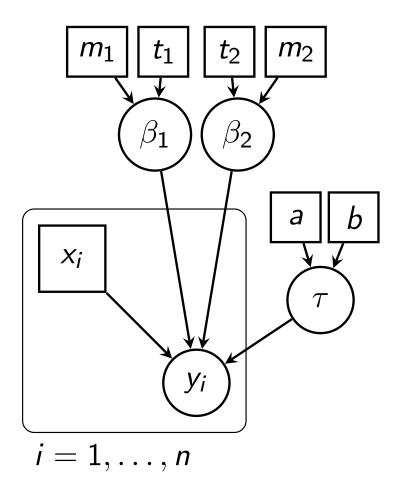
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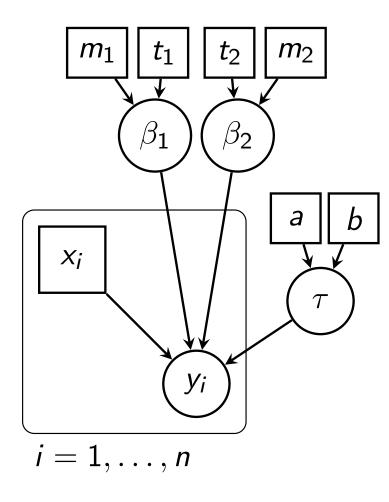
- corresponds to a *fully connected graph*: there is a link between every pair of vertices (each of the K vertices has incoming edges from all lower-numbered vertices)
- sparser graph => nodes have fewer parents
 less complex joint





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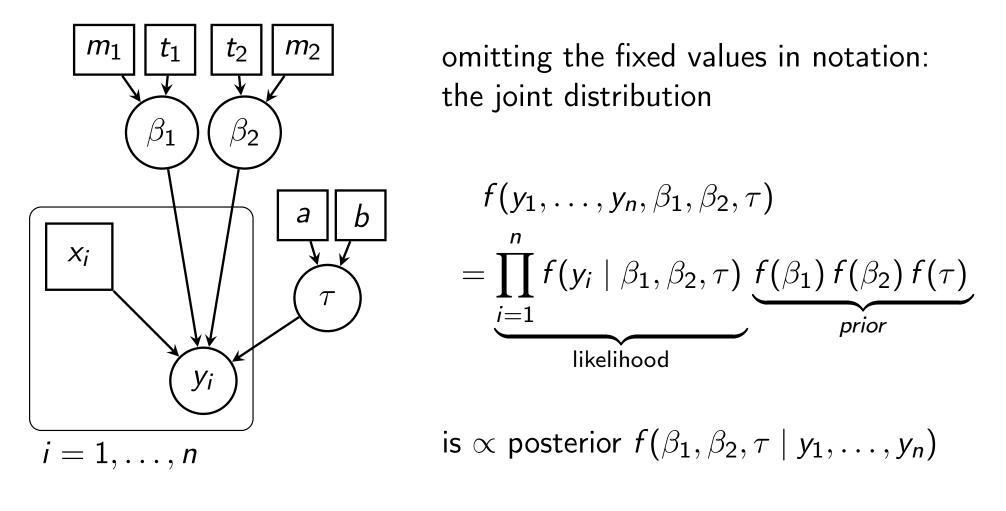
$$f(y_1,\ldots,y_n,\beta_1,\beta_2,\tau) = \prod_{i=1}^n f(y_i \mid \beta_1,\beta_2,\tau) f(\beta_1) f(\beta_2) f(\tau)$$



omitting the fixed values in notation: the joint distribution

$$f(y_1, \dots, y_n, \beta_1, \beta_2, \tau) = \prod_{\substack{i=1 \\ i \neq 1}}^n f(y_i \mid \beta_1, \beta_2, \tau) \underbrace{f(\beta_1) f(\beta_2) f(\tau)}_{prior}$$

is
$$\propto$$
 posterior $f(\beta_1, \beta_2, \tau \mid y_1, \dots, y_n)$



▶ it would be really useful to get posterior estimates based on the non-normalized density $f(y_1, \ldots, y_n, \beta_1, \beta_2, \tau)$!

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 - conditioning = use only samples m with the right value of the conditioning parameter(s)

(or redo the sampling with fixed conditioned values)

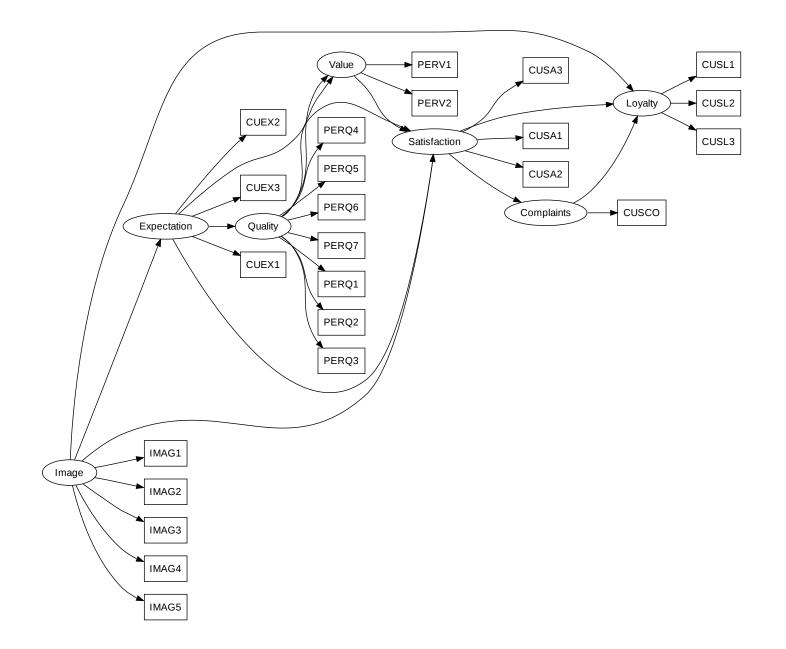
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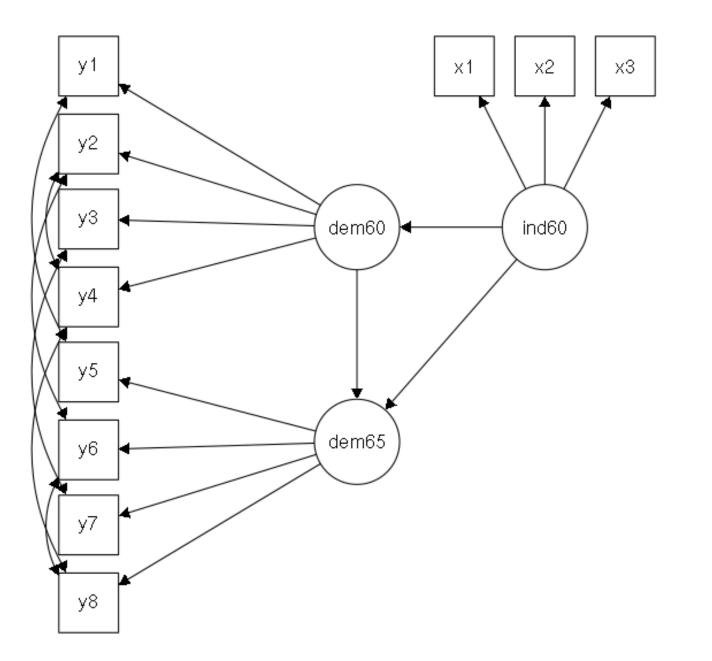
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- here: do sensitivity analysis by varying prior distributions in sets: f(β₁) ∈ Mβ₁,...

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- partial least squares (R package semPLS): iterative fitting of latent variable values and regression coefficients via least squares
- path analysis: special case where a measurement can be linked to only one construct

Bayesian hierarchical modelling, simulation and MCMC _{Outline}

Bayesian hierarchical modelling / Bayesian networks / graphical models

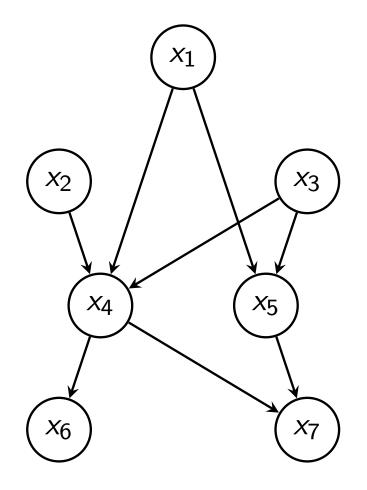
Exercises I

Simulation & MCMC

Exercises II

Exercise 1: Factorization of a Joint

Which factorization of $f({x_i}_{i \in [1,...,7]})$ does this graph encode?



The naive Bayes classifier from Part 6 assumes that the joint distribution of class c and attributes a_1, \ldots, a_k can be factorized as

$$p(c,a) = p(c)p(a \mid c) = p(c)\prod_{i=1}^{k} p(a_i \mid c).$$

Draw the corresponding DAG! (Hint: use either a plate or consider two attributes a_1 and a_2 only.)

Exercise 3: Naive Bayes Classifier with Dirichlet Priors We can introduce parameters for p(c) and $p(a_i | c)$:

$$(n(c))_{c \in C} \sim \text{Multinomal}(\theta_c; c \in C)$$
(36)
$$\forall c \in C \colon (n(a_i, c))_{a_i \in A_i} \mid c \sim \text{Multinomal}(\theta_{a_i \mid c}; a_i \in A_i)$$
(37)

where C denotes the set of all possible class values, and A_i denotes the set of all possible values of attribute *i*.

The θ parameters can be estimated using a Dirichlet prior:

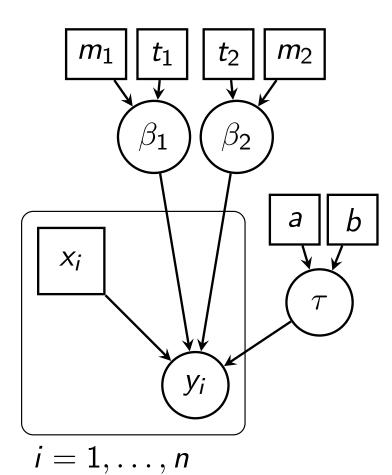
$$(\theta_c)_{c\in\mathcal{C}}\sim \mathsf{Dir}(s,(t(c))_{c\in\mathcal{C}})$$
 (38)

$$\forall c \in \mathcal{C} : (\theta_{a_i|c})_{a_i \in \mathcal{A}_i} \mid c \sim \mathsf{Dir}(s, (t(a_i, c))_{a_i \in \mathcal{A}_i})$$
(39)

where we must have that $\sum_{a_i \in A_i} t(a_i, c) = t(c)$. [Note that t(c) is the prior expectation of θ_c and $t(a_i, c)/t(c)$ is the prior expectation of $\theta_{a_i|c}$.]

Draw the corresponding graph!

Exercise 4: Sensitivity Analysis



In the linear regression example there are 6 hyperparameters m_1, t_1, m_2, t_2, a, b .

How would you do sensitivity analysis over the prior in that example? What problems do you foresee? Bayesian hierarchical modelling, simulation and MCMC _{Outline}

Bayesian hierarchical modelling / Bayesian networks / graphical models

Exercises I

Simulation & MCMC

Exercises II

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- then: MCMC for sampling from multivariate distributions

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$$\lim_{M \to \infty} E(\widehat{g(X)}) \xrightarrow{a.s.} E(g(X)) \text{ (strong law of large numbers)}$$

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• estimate
$$E(g(X))$$
 by $E(\widehat{g(X)}) = \frac{1}{M} \sum_{i=1}^{M} g(x_i)$

• unbiased:
$$E\left(E\left(\widehat{g(X)}\right)\right) = E\left(g(X)\right)$$

• variance: Var
$$\left(E \widehat{(g(X))} \right) = \frac{1}{M} \operatorname{Var} (g(X))$$

for independent samples only!

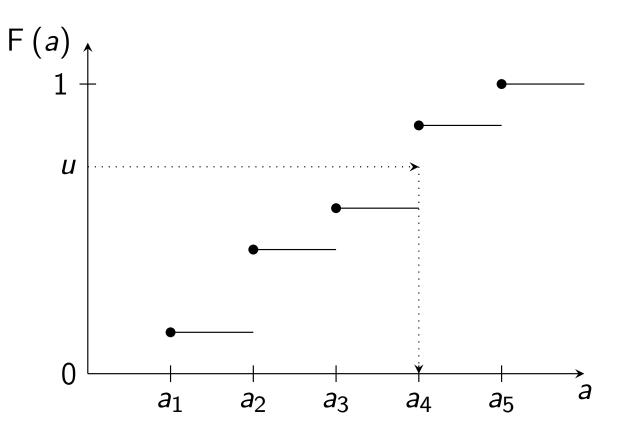
 precision of MC estimate increases with *M*, independent of parameter dimension! (numeric integration: number of evaluation points increases exponentially with dimension)

$$\lim_{M \to \infty} E(\widehat{g(X)}) \xrightarrow{a.s.} E(g(X)) \text{ (strong law of large numbers)}$$

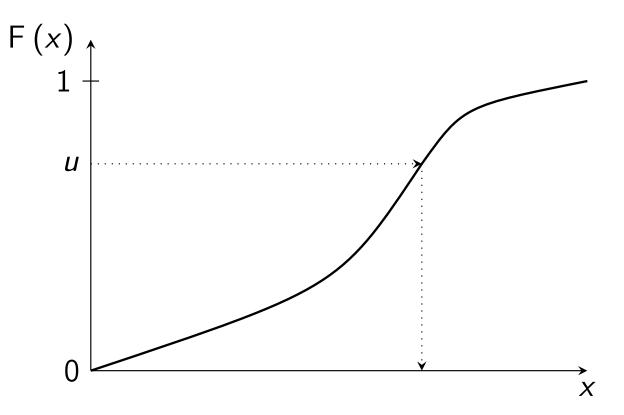
• $E(\widehat{g(X)}) \stackrel{a.s.}{\sim} N(E(g(X)), \frac{1}{M}Var(g(X)))$ (central limit thm)

- assumption for all sampling algorithms:
 we can sample from the uniform U([0, 1])
- ► done by pseudo-random number generator (PNRG), in **R**: ?RNG

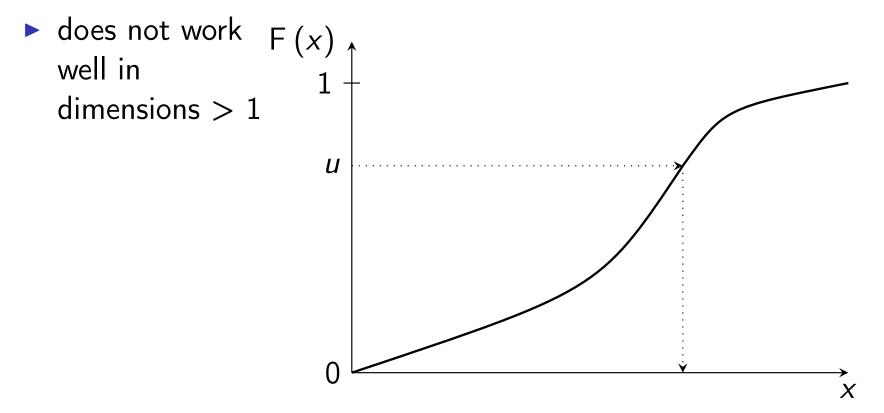
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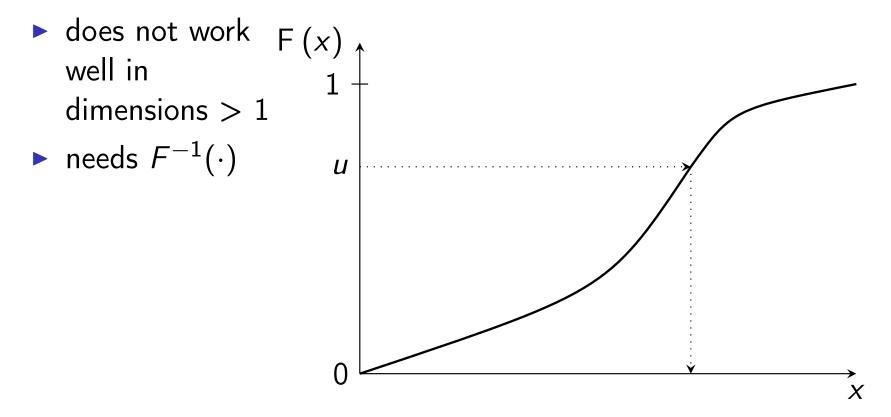
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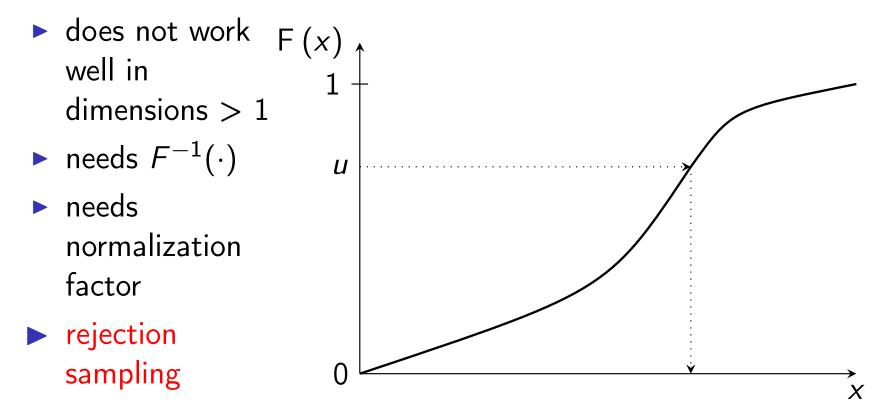


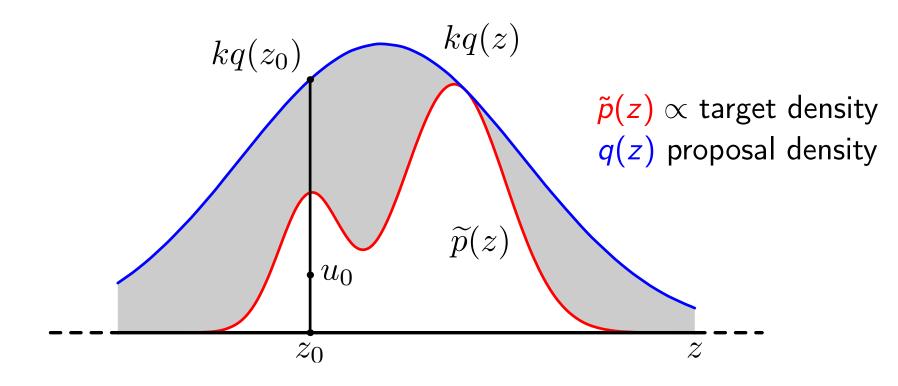
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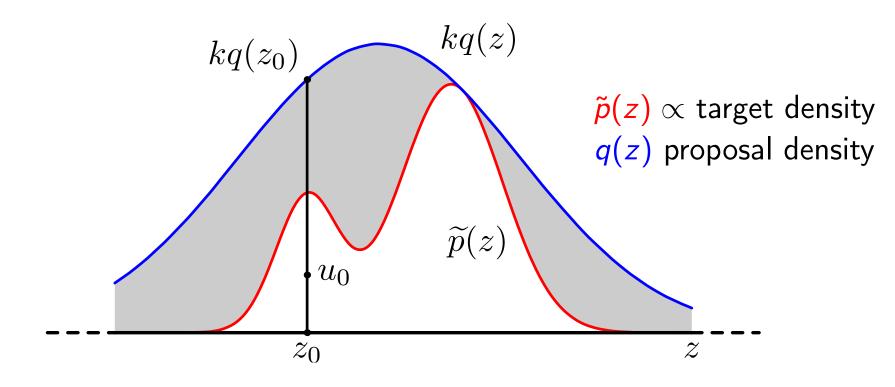


Simulation & MCMC: Univariate Sampling

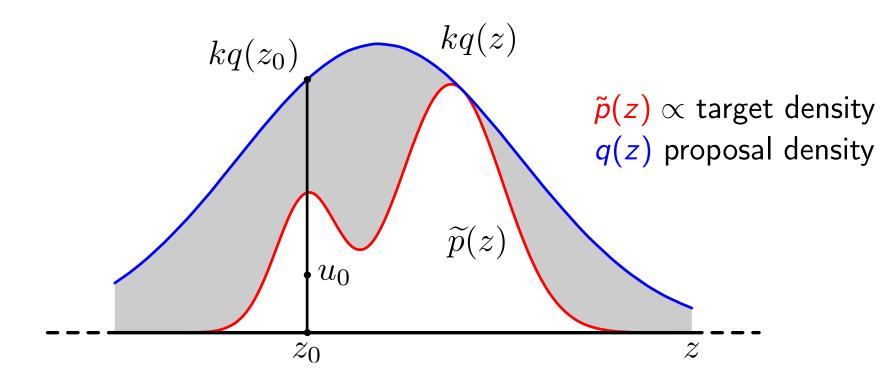
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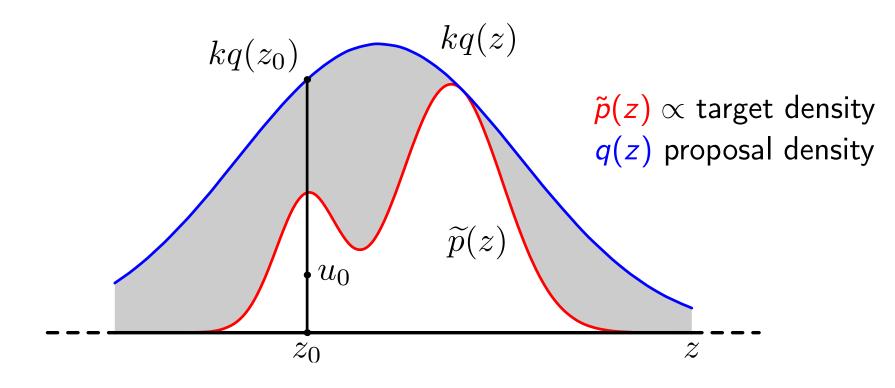




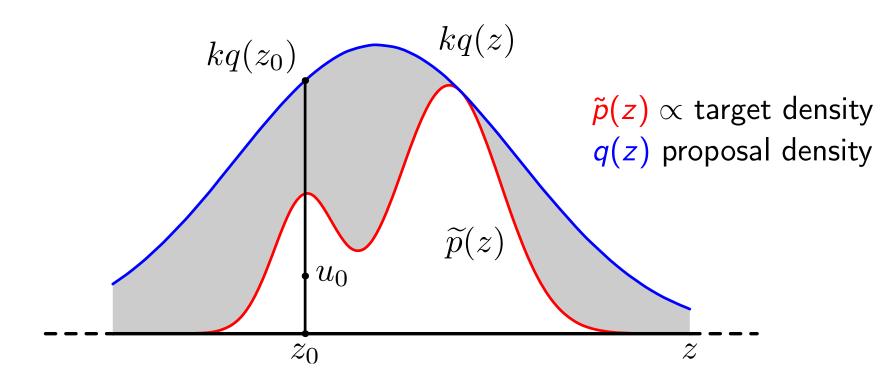
1. sample $z \iff$ from q(z)



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- 2. sample $u(\uparrow)$ from U([0, kq(z)])



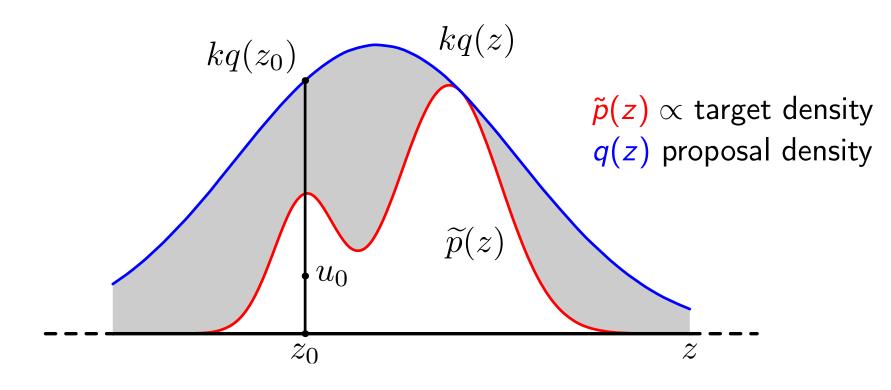
1. sample $z \iff$ from q(z)2. sample $u (\updownarrow)$ from U([0, kq(z)]) sample points uniformly from union of white and grey areas



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sample points uniformly from union of white and grey areas

3. reject all points in the grey area



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sample points uniformly from union of white and grey areas

- 2. sample $u(\uparrow)$ from U([0, kq(z)])
- 3. reject all points in the grey area
- 4. forget about *u*: *z* distributed $\propto \tilde{p}(z)!$

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- idea: produce samples by a Markov Chain: random walk over parameter space

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- but: samples are not independent!

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 propose a step (draw from easy-to-sample-from proposal distribution)

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• loop over parameter vector $(\theta_1, \theta_2, \ldots)$

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- draw from the *full conditionals* $f(\theta_i \mid \text{everything else}) \propto \text{joint}$
- special case of MH where proposals are always accepted

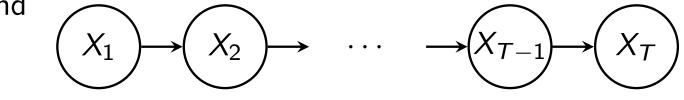
Algorithms create a

- stationary,
- irreducible and
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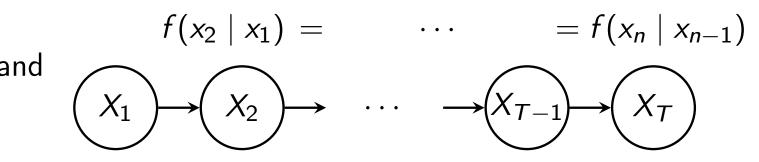
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Markov chain



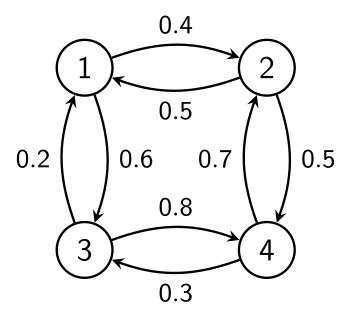
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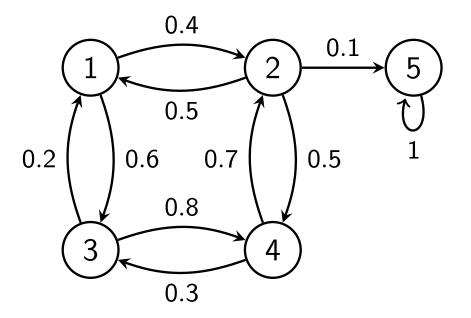
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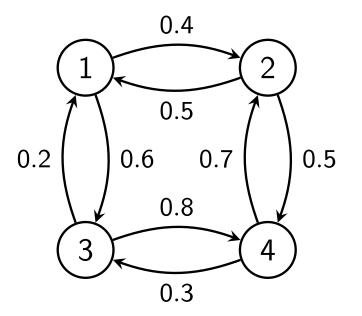
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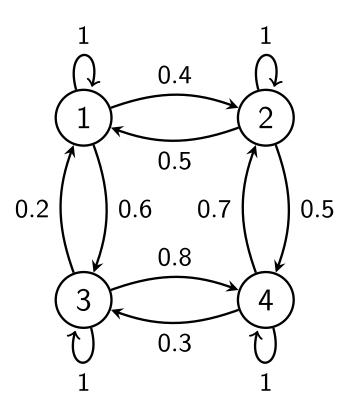
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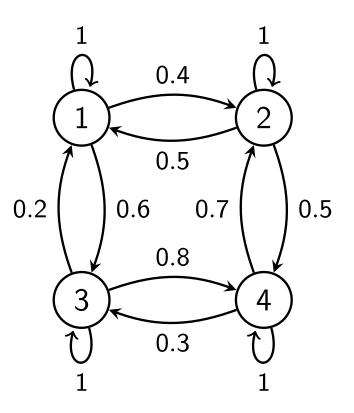
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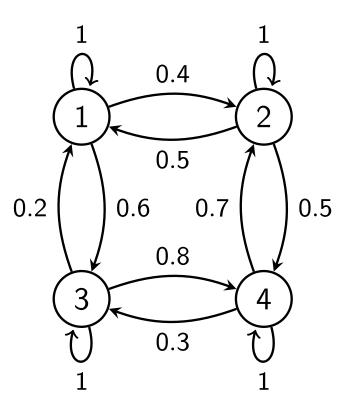
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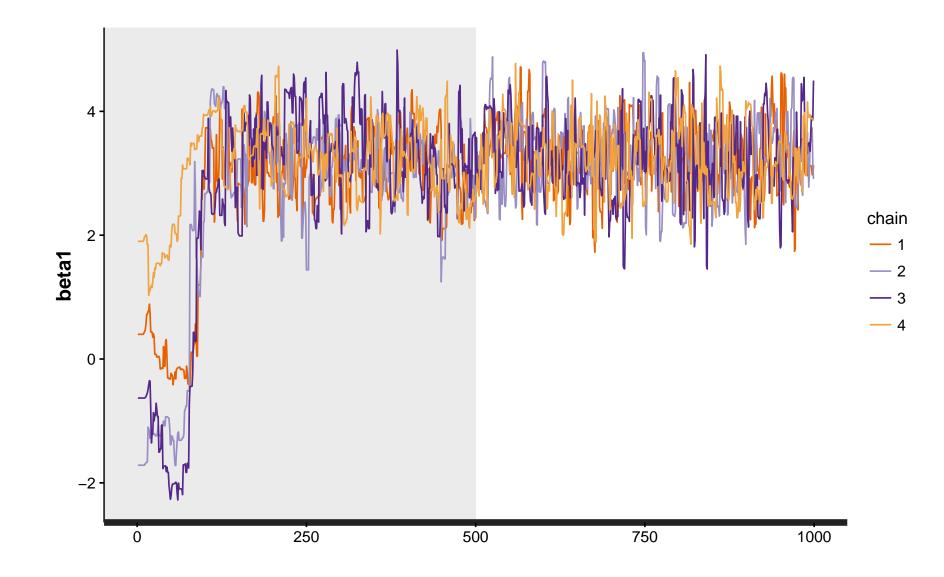
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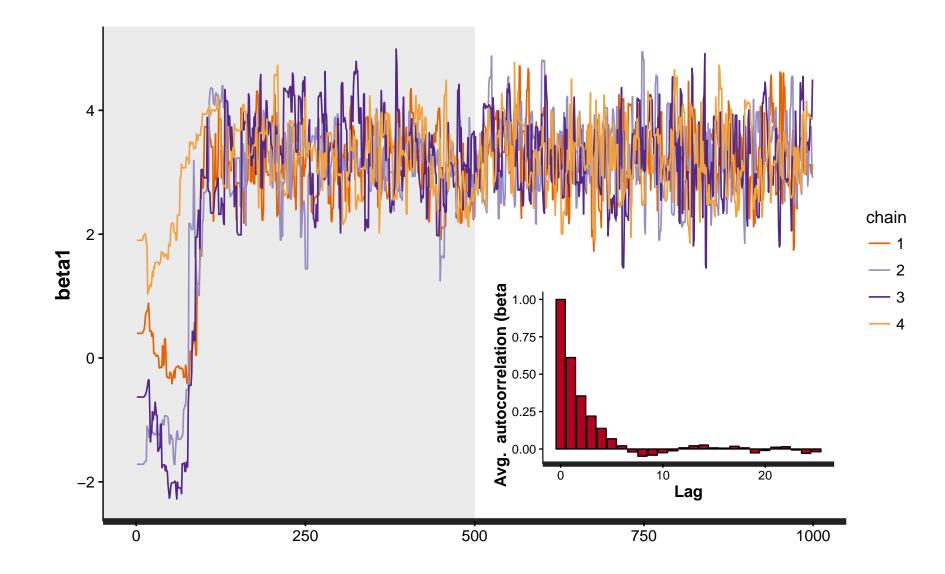


$$(p_1(t), p_2(t), p_3(t), p_4(t)) \xrightarrow{t \to \infty} (p_1, p_2, p_3, p_4)$$

MCMC: Warm-Up (= Burn-In), Mixing, Thinning



MCMC: Warm-Up (= Burn-In), Mixing, Thinning



Bayesian hierarchical modelling, simulation and MCMC _{Outline}

Bayesian hierarchical modelling / Bayesian networks / graphical models

Exercises I

Simulation & MCMC

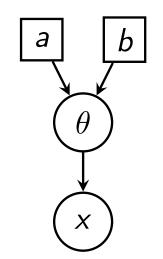
Exercises II

A Stan model is defined by five program blocks:

```
model1 <- "
data {
 • • •
}
transformed data {
 • • •
}
parameters {
 • • •
}
transformed parameters {
 • • •
}
model {
 • • •
}
generated quantities {
  • • •
ጉ"
```

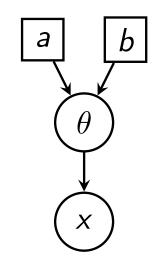
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model1 <- "
data {
  • • •
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    ...
    required
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}
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ጉ"
```



 $f(x \mid \theta) \sim \text{Binomial}(n, \theta)$ $f(\theta \mid a, b) \sim \text{Beta}(a, b)$

```
library(rstan)
model0 <- "
data {
  int <lower=0> n;
  int<lower=0> x;
}
parameters {
  real<lower=0,upper=1> theta;
}
model {
  theta \sim beta(2,2);
  x \sim binomial(n, theta);
}
11
data0 <- list(n=10, x=5)</pre>
```

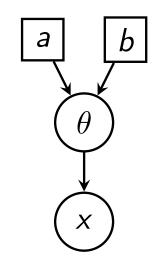


 $f(x \mid \theta) \sim \text{Binomial}(n, \theta)$ $f(\theta \mid a, b) \sim \text{Beta}(a, b)$

```
library(rstan)
model0 <- "
data {
  int<lower=0> n;
  int<lower=0> x;
}
parameters {
  real < lower = 0, upper = 1 > theta;
}
model {
  theta \sim beta(2,2);
  x \sim binomial(n, theta);
}
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data0 <- list(n=10, x=5)
```

Running the model creates a stanfit object.

```
fit0 <- stan(model_code=model0, data=data0, iter=1000, chains=4)
print(fit0); plot(fit0)</pre>
```



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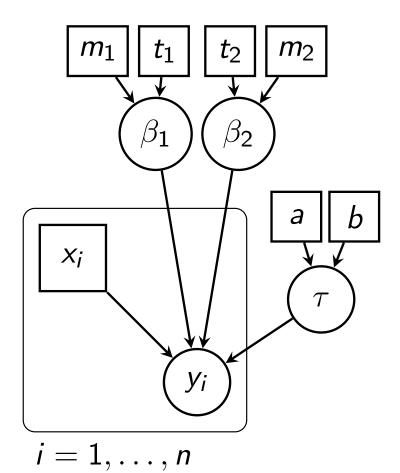
```
library(rstan)
modelO <- "
data {
  int < lower = 0 > n;
  int<lower=0> x;
}
parameters {
  real < lower = 0, upper = 1 > theta;
}
model {
  theta \sim beta(2,2);
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data0 <- list(n=10, x=5)
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```

The samples can be extracted by samples0 = extract(fit0, c("theta")) http://mc-stan.org/documentation/

http://github.com/stan-dev/rstan/wiki/RStan-Getting-Started#how-to-use-rstan



We want to estimate the parameters in the linear regression example, using RStan to sample from the posterior.

The model assumptions are:

 $\begin{array}{ll} y_i \mid \beta_1, \beta_2, \tau \sim \mathsf{N}(\beta_1 + x_i \beta_2, 1/\tau) \\ \tau \mid a, b \sim \mathsf{Gamma}(a, b), & a = b = 10^{-3} \\ \beta_1 \mid m_1, t_1 \sim \mathsf{N}(m_1, 1/t_1), & m_1 = 0, t_1 = 10^4 \\ \beta_2 \mid m_2, t_2 \sim \mathsf{N}(m_2, 1/t_2), & m_2 = 0, t_2 = 10^4 \end{array}$

```
• Create an artificial data set x_1, \ldots, x_n, y_1, \ldots, y_n by
```

```
data <- list()
data$N <- 50
data$x <- rnorm(data$N)+30
data$y <- 3 + 5*data$x + rnorm(data$N, sd=1/10)</pre>
```

What are thus the 'true' parameter values?

- Define the model in Stan. Include a transformed parameters block where you define $\sigma = \sqrt{1/\tau}$. (In Stan, the Normal distribution is parametrized with the standard deviation σ !)
- Simulate four chains with 1000 iterations each and use plot() and print() to get a first impression of the results. What point estimates do you get for β₁, β₂ and σ?

- The functions stan_trace(), stan_dens() and stan_ac() allow you to analyze your sample from the posterior distribution more closely. (You can include the warm-up phase in your plots by setting inc_warmup = TRUE.) How long is the warm-up phase? Do your chains mix well? Is thinning necessary?
- The function pairs() also works on stanfit objects.
 Plot pairwise scatterplots of your sample using pairs().
 What do you observe about the relation between β₁ and β₂?

The high correlation between β₁ and β₂ indicates that the Markov chain cannot move around freely. You can mitigate this problem by centering the data x₁,..., x_n. The mean for the Normal distribution of y_i is then given by β₁^c + β₂(x_i - x̄), where β₁^c = β₁ + β₂x̄. Add the following block to your stan model definition,

```
transformed data {
   vector[N] xcentered;
   xcentered=x-mean(x);
}
```

and edit the parameters and model blocks such that the model generates samples from β_1^c instead of β_1 .

- Edit the transformed parameters block to define β_1 as $\beta_1 = \beta_1^c + \beta_2 \bar{x}$.
- Simulate four chains with 1000 iterations each from this new model, and analyze your sample from the posterior distribution like for the first model. What has changed?

Choose an informative prior for one or both of β₁ and β₂.
 Try out different values for mean and standard deviation.
 What is the effect on the chains and the posterior densities?