Wednesday 14:00-17:30

## Part 7

# Bayesian hierarchical modelling, simulation and MCMC 

by Gero Walter

# Bayesian hierarchical modelling, simulation and MCMC 

 OutlineBayesian hierarchical modelling / Bayesian networks / graphical models

## Exercises I

Simulation \& MCMC

Exercises II

## Bayesian Hierarchical Modelling, a.k.a. Bayesian (Belief)

 Networks, a.k.a. Graphical Models- many names for the same thing (it's a powerful tool), I will use the term Bayesian Networks (BNs)
- BNs as a unifying way to think about (Bayesian) statistical models
- how to build complex Bayesian models out of simple building blocks
- how to specify joint distributions (over many variables) via univariate distributions using conditional independence assumptions
- conditional independence assumptions are visualized by a graph
- the graph can establish a hierarchy between variables


## Bayesian Networks: Simple Example



## Another Example: Linear Regression



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$$
\begin{aligned}
& y_{i} \mid \mu_{i}, \tau \sim \mathrm{~N}\left(\mu_{i}, 1 / \tau\right), \quad \text { where } \mu_{i}=\beta_{1}+x_{i} \beta_{2} \\
& \text { - } y_{i} \mid \beta_{1}, \beta_{2}, \tau \sim \mathrm{~N}\left(\beta_{1}+x_{i} \beta_{2}, 1 / \tau\right) \\
& \tau \mid a, b \sim \operatorname{Gamma}(a, b) \\
& a=b=10^{-3} \\
& \beta_{1} \mid m_{1}, t_{1} \sim \mathrm{~N}\left(m_{1}, 1 / t_{1}\right) \\
& \beta_{2} \mid m_{2}, t_{2} \sim \mathrm{~N}\left(m_{2}, 1 / t_{2}\right) \\
& m_{1}=m_{2}=0, \quad t_{1}=t_{2}=10^{4}
\end{aligned}
$$

## Another Example: Linear Regression

$$
\begin{aligned}
& \text { (1) } \quad y_{i}=\beta_{1}+x_{i} \beta_{2}+\varepsilon_{i}, \quad \text { where } \varepsilon_{i} \stackrel{t_{2}}{\sim} \\
& \text { 病 } \mid \mu_{i}, \tau \sim \mathrm{~N}\left(\mu_{i}, 1 / \tau\right) \text {, where } \mu_{i}=\beta_{1}+x_{i} \beta_{2} \\
& \rightarrow y_{i} \mid \beta_{1}, \beta_{2}, \tau \sim \mathrm{~N}\left(\beta_{1}+x_{i} \beta_{2}, 1 / \tau\right) \\
& \tau \mid a, b \sim \operatorname{Gamma}(a, b) \\
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graph:
cond. indep. relations
variables:
conditional distributions

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- often: $\mathrm{BN}=$ discrete variables only, distributions defined via conditional probability tables (CPTs)
- What kind of graphs work for expressing conditional independence relations?


## Bayesian Networks: Directed Acyclic Graphs

Definition (Directed Graph)
A directed graph $G=(V, E)$ consists of a set of vertices $V$ and a set of edges $E$, where $E \subset V \times V$. An arrow leads from $u \in V$ to $v \in V$ if and only if $(u, v) \in E ; u$ is the source and $v$ is the target of edge ( $u, v$ ).

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Definition (Paths and Cycles)
A path in a graph is an ordered set of edges $\left\{e_{i}\right\}$ such that $t\left(e_{i}\right)=s\left(e_{i+1}\right)$ (chain of head-to-tail arrows). A cycle is a path such that $t\left(e_{N}\right)=s\left(e_{1}\right)$, where $N$ is the number of edges in the path.

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## Definition (Directed Acyclic Graph)

A directed acyclic graph (DAG) is a directed graph that does not contain a cycle, i.e. there does not exist a subset of edges that forms a cycle.

## Bayesian Networks: Formal Definition

## Definition (Bayesian Network)

Given a DAG $G=(V, E)$, and variables $x_{V}=\left\{x_{v}\right\}_{v \in V}$, a Bayesian network with respect to $G$ and $x_{V}$ is a joint probability distribution for the $x_{V}$ of the form:

$$
f\left(x_{V}\right)=\prod_{v \in V} f\left(x_{v} \mid x_{\mathrm{pa}(v)}\right)
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where $\mathrm{pa}(v)$ is the set of parents of $v$, i.e. the set of vertices $u$ such that $(u, v)$ is an edge.

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- joint distribution factorizes according to the graph!


## Bayesian Networks: Factorization of the joint

One can always factorize a joint distribution by

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\begin{array}{rl}
f\left(x_{1}, \ldots, x_{K}\right)=f & f\left(x_{K} \mid x_{1}, \ldots, x_{K-1}\right) f\left(x_{K-1} \mid x_{1}, \ldots, x_{K-2}\right) \\
& \cdots f\left(x_{3} \mid x_{1}, x_{2}\right) f\left(x_{2} \mid x_{1}\right) f\left(x_{1}\right)
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- corresponds to a fully connected graph: there is a link between every pair of vertices (each of the $K$ vertices has incoming edges from all lower-numbered vertices)
- sparser graph $\Longrightarrow$ nodes have fewer parents
$\Longrightarrow$ less complex joint

Factorization of the joint: Example


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omitting the fixed values in notation: the joint distribution

$$
\begin{aligned}
& f\left(y_{1}, \ldots, y_{n}, \beta_{1}, \beta_{2}, \tau\right) \\
= & \prod_{i=1}^{n} f\left(y_{i} \mid \beta_{1}, \beta_{2}, \tau\right) f\left(\beta_{1}\right) f\left(\beta_{2}\right) f(\tau)
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- it would be really useful to get posterior estimates based on the non-normalized density $f\left(y_{1}, \ldots, y_{n}, \beta_{1}, \beta_{2}, \tau\right)$ !


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- marginalizing $=$ use only, e.g., $\left\{\beta_{1}^{m}\right\}, m=1, \ldots, M$
- conditioning = use only samples $m$ with the right value of the conditioning parameter(s) (or redo the sampling with fixed conditioned values)


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- conditional independence with IP gets very non-trivial (see, e.g., $[2, \S 4]$ for the gory details)


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- use sets of conditional distributions at nodes: credal networks (see, e.g., [2, §10], [17])
- specific algorithms for discrete credal networks (see, e.g., [2, §10.5.3], or [14])
- conditional independence with IP gets very non-trivial (see, e.g., $[2, \S 4]$ for the gory details)
- here: do sensitivity analysis by varying prior distributions in sets: $f\left(\beta_{1}\right) \in \mathcal{M}_{\beta_{1}}, \ldots$


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- path analysis: special case where a measurement can be linked to only one construct


# Bayesian hierarchical modelling, simulation and MCMC 

 OutlineBayesian hierarchical modelling / Bayesian networks / graphical models

Exercises I

Simulation \& MCMC

Exercises II

## Exercise 1: Factorization of a Joint

Which factorization of $f\left(\left\{x_{i}\right\}_{i \in[1, \ldots, 7]}\right)$
does this graph encode?


## Exercise 2: Naive Bayes Classifier

The naive Bayes classifier from Part 6 assumes that the joint distribution of class $c$ and attributes $a_{1}, \ldots, a_{k}$ can be factorized as

$$
p(c, a)=p(c) p(a \mid c)=p(c) \prod_{i=1}^{k} p\left(a_{i} \mid c\right)
$$

Draw the corresponding DAG!
(Hint: use either a plate or consider two attributes $a_{1}$ and $a_{2}$ only.)

## Exercise 3: Naive Bayes Classifier with Dirichlet Priors

We can introduce parameters for $p(c)$ and $p\left(a_{i} \mid c\right)$ :

$$
\begin{align*}
(n(c))_{c \in \mathcal{C}} & \sim \operatorname{Multinomal}\left(\theta_{c} ; c \in \mathcal{C}\right)  \tag{36}\\
\forall c \in \mathcal{C}:\left(n\left(a_{i}, c\right)\right)_{a_{i} \in \mathcal{A}_{i}} \mid c & \sim \operatorname{Multinomal}\left(\theta_{a_{i} \mid} \mid c ; a_{i} \in \mathcal{A}_{i}\right) \tag{37}
\end{align*}
$$

where $\mathcal{C}$ denotes the set of all possible class values, and $\mathcal{A}_{i}$ denotes the set of all possible values of attribute $i$.

The $\theta$ parameters can be estimated using a Dirichlet prior:

$$
\begin{align*}
\left(\theta_{c}\right)_{c \in \mathcal{C}} & \sim \operatorname{Dir}\left(s,(t(c))_{c \in \mathcal{C}}\right)  \tag{38}\\
\forall c \in \mathcal{C}:\left(\theta_{a_{i} \mid c}\right)_{a_{i} \in \mathcal{A}_{i}} \mid c & \sim \operatorname{Dir}\left(s,\left(t\left(a_{i}, c\right)\right)_{a_{i} \in \mathcal{A}_{i}}\right) \tag{39}
\end{align*}
$$

where we must have that $\sum_{a_{i} \in \mathcal{A}_{i}} t\left(a_{i}, c\right)=t(c)$. [Note that $t(c)$ is the prior expectation of $\theta_{c}$ and $t\left(a_{i}, c\right) / t(c)$ is the prior expectation of $\theta_{a_{i} \mid c}$.]

Draw the corresponding graph!

## Exercise 4: Sensitivity Analysis



In the linear regression example there are 6 hyperparameters $m_{1}, t_{1}, m_{2}, t_{2}, a, b$.

How would you do sensitivity analysis over the prior in that example? What problems do you foresee?

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- ... any function of posterior parameters by sample equivalent
- first: quick look at sampling from univariate distributions
- then: MCMC for sampling from multivariate distributions

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- variance: $\operatorname{Var}(\mathrm{E} \widehat{(g(X)}))=\frac{1}{M} \operatorname{Var}(g(X))$ for independent samples only!


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$\left.-\lim _{M \rightarrow \infty} \mathrm{E} \widehat{(g(X)}\right) \xrightarrow{\text { a.s. }} \mathrm{E}(g(X))$ (strong law of large numbers)


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- variance: $\operatorname{Var}(\mathrm{E} \widehat{(g(X))})=\frac{1}{M} \operatorname{Var}(g(X))$ for independent samples only!
- precision of MC estimate increases with $M$, independent of parameter dimension! (numeric integration: number of evaluation points increases exponentially with dimension)
- $\left.\lim _{M \rightarrow \infty} \mathrm{E} \widehat{(g(X)}\right) \xrightarrow{\text { a.s. }} \mathrm{E}(g(X))$ (strong law of large numbers)
- $\mathrm{E} \widehat{(g(X))} \stackrel{\text { a.s. }}{\sim} \mathrm{N}\left(\mathrm{E}(g(X)), \frac{1}{M} \operatorname{Var}(g(X))\right)$ (central limit thm)


## Simulation \& MCMC: Univariate Sampling

- assumption for all sampling algorithms: we can sample from the uniform $U([0,1])$
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- does not work $\mathrm{F}(x)$ well in dimensions $>1$
- needs $F^{-1}(\cdot)$
- needs normalization factor
- rejection sampling



## Simulation \& MCMC: Rejection Sampling



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4. forget about $u: z$ distributed $\propto \tilde{p}(z)$ !

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- but: samples are not independent!

Markov Chain Monte Carlo: Algorithms

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- special case of MH where proposals are always accepted


# Markov Chain Monte Carlo: Why does this work? 

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$$
\left(p_{1}(t), p_{2}(t), p_{3}(t), p_{4}(t)\right) \xrightarrow{t \rightarrow \infty}\left(p_{1}, p_{2}, p_{3}, p_{4}\right)
$$

MCMC: Warm-Up ( = Burn-In), Mixing, Thinning


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# Bayesian hierarchical modelling, simulation and MCMC 

 OutlineBayesian hierarchical modelling / Bayesian networks / graphical models

## Exercises |

Simulation \& MCMC

Exercises II

## Exercise: Quick start RStan

A Stan model is defined by five program blocks:

```
model1 <- "
data {
}
transformed data {
}
parameters {
}
transformed parameters {
}
model {
}
generated quantities {
}"
```


## Exercise: Quick start RStan

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$f(x \mid \theta) \sim \operatorname{Binomial}(n, \theta)$
$f(\theta \mid a, b) \sim \operatorname{Beta}(a, b)$
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```
library (rstan)
model0 <- "
data \{
    int<lower=0> n;
    int<lower=0> \(x\);
\}
parameters \{
    real<lower=0, upper=1> theta;
\}
model \{
    theta \(\sim\) beta \((2,2)\);
    \(\mathrm{x} \sim\) binomial (n, theta);
\}
"
data0 <- list \((\mathrm{n}=10, \mathrm{x}=5)\)
```


## Exercise: Quick start RStan



```
```

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model0 <- "
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parameters {
parameters {
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}
model {
model {
theta ~ beta(2,2);
theta ~ beta(2,2);
x ~ binomial(n, theta);
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data0 <- list(n=10, x=5)

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$f(x \mid \theta) \sim \operatorname{Binomial}(n, \theta)$
$f(\theta \mid a, b) \sim \operatorname{Beta}(a, b)$

Running the model creates a stanfit object.
fit0 <- stan (model_code=model0, data=data0, iter=1000, chains=4) print (fito) ; plot(fit0)

## Exercise: Quick start RStan


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```
library(rstan)
model0 <- "
data {
    int<lower=0> n;
    int<lower=0> x;
}
parameters {
    real<lower=0,upper=1> theta;
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model {
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    x ~ binomial(n, theta);
}
"
data0 <- list(n=10, x=5)
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Running the model creates a stanfit object.
fit0 <- stan (model_code=model0, data=data0, iter=1000, chains=4) print (fito) ; plot (fit0)

The samples can be extracted by samples0 $=$ extract(fit0, c("theta")) http://mc-stan.org/documentation/

## Exercise: Linear regression in RStan



We want to estimate the parameters in the linear regression example, using RStan to sample from the posterior.

The model assumptions are:
$y_{i} \mid \beta_{1}, \beta_{2}, \tau \sim \mathrm{~N}\left(\beta_{1}+x_{i} \beta_{2}, 1 / \tau\right)$
$\tau \mid a, b \sim \operatorname{Gamma}(a, b), \quad a=b=10^{-3}$
$\beta_{1} \mid m_{1}, t_{1} \sim \mathrm{~N}\left(m_{1}, 1 / t_{1}\right), \quad m_{1}=0, t_{1}=10^{4}$
$\beta_{2} \mid m_{2}, t_{2} \sim N\left(m_{2}, 1 / t_{2}\right), \quad m_{2}=0, t_{2}=10^{4}$

## Exercise: Linear regression in RStan

- Create an artificial data set $x_{1}, \ldots, x_{n}, y_{1}, \ldots, y_{n}$ by

```
data <- list()
data$N <- 50
data$x <- rnorm(data$N)+30
data$y <- 3 + 5*data$x + rnorm(data$N, sd=1/10)
```

What are thus the 'true' parameter values?

- Define the model in Stan. Include a transformed parameters block where you define $\sigma=\sqrt{1 / \tau}$. (In Stan, the Normal distribution is parametrized with the standard deviation $\sigma$ !)
- Simulate four chains with 1000 iterations each and use plot() and print() to get a first impression of the results. What point estimates do you get for $\beta_{1}, \beta_{2}$ and $\sigma$ ?


## Exercise: Linear regression in RStan

- The functions stan_trace(), stan_dens() and stan_ac() allow you to analyze your sample from the posterior distribution more closely. (You can include the warm-up phase in your plots by setting inc_warmup = TRUE.) How long is the warm-up phase? Do your chains mix well? Is thinning necessary?
- The function pairs() also works on stanfit objects. Plot pairwise scatterplots of your sample using pairs(). What do you observe about the relation between $\beta_{1}$ and $\beta_{2}$ ?


## Exercise: Linear regression in RStan

- The high correlation between $\beta_{1}$ and $\beta_{2}$ indicates that the Markov chain cannot move around freely. You can mitigate this problem by centering the data $x_{1}, \ldots, x_{n}$. The mean for the Normal distribution of $y_{i}$ is then given by $\beta_{1}^{c}+\beta_{2}\left(x_{i}-\bar{x}\right)$, where $\beta_{1}^{c}=\beta_{1}+\beta_{2} \bar{x}$.
Add the following block to your stan model definition,

```
transformed data {
    vector[N] xcentered;
    xcentered=x-mean(x);
}
```

and edit the parameters and model blocks such that the model generates samples from $\beta_{1}^{c}$ instead of $\beta_{1}$.

- Edit the transformed parameters block to define $\beta_{1}$ as $\beta_{1}=\beta_{1}^{c}+\beta_{2} \bar{x}$.
- Simulate four chains with 1000 iterations each from this new model, and analyze your sample from the posterior distribution like for the first model. What has changed?


## Exercise: Linear regression in RStan

- Choose an informative prior for one or both of $\beta_{1}$ and $\beta_{2}$. Try out different values for mean and standard deviation. What is the effect on the chains and the posterior densities?

